Oncology

The association between interleukin-1 β gene polymorphisms and the risk of breast cancer: a systematic review and meta-analysis

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Abstract

Introduction: It is reported that there is a close association between interleukin-1 β (IL-1 β) gene polymorphisms and breast cancer risk. However, the results remain controversial.

Material and methods: Eligible published articles were searched in PubMed, Embase, and Web of Science databases up to June 2018. Odds ratios with 95% confidence intervals were used to identify potential links between IL-1 β genetic polymorphisms and the risk of breast cancer.

Results: From our results, we found that three common polymorphisms in IL-1 β (rs16944, rs1143634, rs1143627) had no significant associations with breast cancer risk in all genetic models. Based on the analysis from ethnic subgroups, there was a higher risk of breast cancer for rs16944 polymorphism in the recessive model and heterozygous model among Asians (TT vs. CC+CT: 1.229, 95% CI: 1.063-1.422, p = 0.005; TT vs. CT: 1.211, 95% CI: 1.057–1.388, p = 0.006). For the rs1143627 polymorphism, a significantly decreased breast cancer risk was observed in the dominant model only in Asians (CT+TT vs. CC: OR = 0.944, 95% CI: 0.897-0.994, p = 0.027). After stratifying patients according to the menopausal state, we found that polymorphism of rs1143627 correlated with reduced breast cancer risk among post-menopausal women in three genotype models: allele, recessive model and homozygous model (T vs C: 0.859, 95% CI: 0.753–0.98, p = 0.024; TT vs. CC+CT: 0.727, 95% CI: 0.576-0.918, p = 0.007; TT vs. CC: 0.743, 95% CI: 0.626–0.882, p = 0.001). As for other analyses with reference to source of controls and genotyping methods, no significant association between IL-1β polymorphism and breast cancer risk was demonstrated. Conclusions: The rs16944 and rs1143627 polymorphisms are significantly associated with the risk of breast cancer only in Asian people and in postmenopausal women respectively.

Key words: polymorphism, breast cancer, interleukin-1β, meta-analysis.

Introduction

Breast cancer is a complex multiple process influenced by multiple factors. It is considered that specific gene polymorphisms have effects on gene transcription, mRNA stability and protein activity [1]. Multi-functional cytokines involved in the process of inflammatory and

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Creative Commons licenses: This is an Open Access article distributed under the terms of the Creative Commons Attribution-NonCommercial-ShareAlike 4.0 International (CC BY -NC -SA 4.0). License (http://creativecommons.org/licenses/by-nc-sa/4.0/) immunological responses are closely associated with the pathogenesis of autoimmune and malignant diseases, making them potential risks for breast cancer [2–4]. The interleukin 1 gene family located on chromosome 2q14.2 includes three members: IL-1 α , IL-1 β and IL-1 receptor antagonist (IL-RA) encoded by IL-1RN [5]. Interleukin 1 α and IL-1 β are potent proinflammatory cytokines, whereas IL-RA is an anti-inflammatory cytokine [6]. Interleukin 1 β can be produced by various cells and it modifies the process of host response to microbial invasion, tissue injury and inflammation [7].

The interleukin 1 β gene has three potentially functional SNPs: -31 (rs1143627, C>T), -511 (rs16944, C>T) in the promoter region and +3954 (rs1143634, C>T) in exon 5 [8, 9]. So far, many studies have been conducted to assess the relations between the three SNPs (rs16944, rs1143634 and rs1143627) in IL-1 β and breast cancer risk [10–20]. However, the results remain conflicting. Therefore, we performed this meta-analysis in order to obtain a more precise evaluation of these links.

Material and methods

Literature search

Relevant studies published before June 1st, 2018 were identified through a search in PubMed, Embase, and Web of Science using a combination of the following terms: ("polymorphism" or "SNP" or "single nucleotide polymorphisms"), ("breast cancer" or "breast carcinoma" or "breast tumor") and ("Interleukin-1" or "IL-1 β " or "Interleukin-1 beta"). The references from the eligible articles or textbooks were also manually searched by us to find additional potential sources.

Inclusion and exclusion criteria

The criteria for studies in our meta-analysis are as follows: (a) studies concentrated on relations between IL-1 β polymorphisms and breast cancer risk; (b) providing sufficient data for the frequencies of alleles and genotypes; (c) published in English. Studies were excluded when: (a) they were not case-control studies; (b) did not supply complete and essential information; (c) they were meta-analyses, reviews, or editorial articles.

Data extraction

Data were extracted from each publication independently by two authors (Wang and Yuan) based on the inclusion criteria mentioned above. For each study, the data were collected as follows: the first author, year of publication, country of origin, ethnicity, menopausal state, numbers of patients and controls, source of controls, mutation detection methods, genotyping methods, minor allele frequency (MAF), allele and genotype frequencies and the evidence of Hardy-Weinberg equilibrium (HWE) in controls. Disputes were settled by consulting with a third author if disagreements occured.

Statistical analysis

In order to assess the strength of associations between IL-1 β gene polymorphisms and breast cancer susceptibility under five genetic models which include the allele model, dominant model, recessive model, homozygous model and the heterozygous model, crude odds ratios (ORs) with their corresponding 95% confidence intervals (CIs) were adopted [21-24]. Hardy-Weinberg equilibrium in control groups was estimated using the χ^2 test. The pooled OR's statistical significance was verified using the Z test with a two-tailed p < 0.05 which is regarded as statistically significant. Between-study variations and heterogeneities were evaluated using either Cochran's Q-statistic or *I*² test. When the result is a *p*-value < 0.05 or l^2 > 50%, it indicates the existence of heterogeneity among studies; only in this circumstance was the random effects model (DerSimonian Laird method) used. If not, the fixed effects model (Mantel-Haenszel method) was performed. In order to explore sources of heterogeneity, we performed subgroup analysis with reference to ethnicity, menopausal state, source of controls and genotyping methods. Additionally, to investigate the sensitivity, we removed each study in turn to evaluate the guality and consistency of results. We used Begg's funnel plot and Egger's linear regression test to detect publication biases. All analyses were conducted with STATA version 12.0.

Results

Eligible studies

Four studies with controls deviating from HWE were excluded [12, 15, 18, 20]. Ten case-control studies were conducted by us to assess associations between IL-1 β polymorphisms and breast cancer risk. As shown in Tables I and II, 6 studies were eligible for rs16944(C>T) including 2454 cases and 2720 controls [11, 12, 14, 16, 18, 19]. Details are as follows: 1) Caucasians, Asians, and Africans were investigated in 4, 2 and 1 studies respectively, 3 of which were associated with menopause. 2) Only 1 study was based on PB regarding the source of controls. 3) Genotype methods included TaqMan, PCR, sequencing and MALDI-TOF. For rs1143634(C>T), it was the same as above, but only two studies were qualified, which

SN	Ps/First author	Year	Racial descent	Country	Menopausal state	Source of controls	Genotype methods	Cases	Controls
rs1	6944:								
	Smith	2004	Caucasian	UK	Ν	РВ	Taqman	141	261
	Hefler	2005	Caucasian	Germany	Ν	HB	PCR	269	227
	Liu	2006	Asian	China	Ν	HB	PCR	365	631
	Balasubramanian	2006	Caucasian	UK	Ν	HB	PCR	703	489
	Pooja	2012	Asian	India	Pre- menopausal	HB	Sequence	107	200
	Pooja	2012	Asian	India	Post- menopausal	HB	Sequence	93	200
	Gong	2013	Caucasian	American	Pre- menopausal	HB	MALDL-TOF	185	163
	Gong	2013	Caucasian	American	Post- menopausal	HB	MALDL-TOF	141	146
	Gong	2013	African	American	Pre- menopausal	HB	MALDL-TOF	237	195
	Gong	2013	African	American	Post- menopausal	HB	MALDL-TOF	213	208
	Zuo	2018	Asian	China	Ν	HB	Sequenom MassARRAY	530	628
rs1	143634:							·	
	Snoussi	2005	African	Tunisia	Ν	HB	PCR	305	200
	Hefler	2005	Caucasian	Germany	Ν	HB	PCR	269	227
	Balasubramanian	2006	Caucasian	UK	Ν	HB	PCR	691	420
	Pooja	2012	Asian	India	Pre- menopausal	HB	Sequence	107	200
	Pooja	2012	Asian	India	Post- menopausal	HB	Sequence	93	200
	Pooja	2012	Asian	India	Ν	HB	Sequence	200	200
rs1	143627:								
	Ito	2002	Asian	Japan	Ν	HB	PCR-CTPP	227	185
	Liu	2006	Asian	China	Ν	HB	PCR	365	631
	Lee	2006	Asian	Korea	Pre- menopausal	РВ	PCR-CTPP	353	290
	Lee	2006	Asian	Korea	Post- menopausal	РВ	PCR-CTPP	206	215
	Lee	2006	Asian	Korea	Ν	РВ	PCR-CTPP	559	505
	Akisik	2007	Asian	Turkey	Ν	Ν	PCR-RFLP	126	110
	Gong	2013	Caucasian	American	Pre- menopausal	HB	MALDL-TOF	186	166
	Gong	2013	Caucasian	American	Post- menopausal	HB	MALDL-TOF	142	146
	Gong	2013	African	American	Pre- menopausal	HB	MALDL-TOF	239	195
	Gong	2013	African	American	Post- menopausal	HB	MALDL-TOF	216	210
	Zuo	2018	Asian	China	Ν	HB	Sequenom MassARRAY	530	628

SNPs/First author	First author Genetype distribution						HWE test						
			Cas	ses					Con	trols			-
	С	Т	СС	СТ	TT	MAF	С	т	СС	СТ	TT	MAF	-
rs16944:													
Smith	187	95	60	67	14	0.34	309	213	87	135	39	0.41	0.25
Hefler	362	176	124	114	31	0.33	287	167	88	111	28	0.37	0.61
Liu	358	372	94	170	101	0.49	699	563	197	305	129	0.45	0.58
Balasubramanian	972	434	339	294	70	0.31	670	308	232	206	51	0.31	0.61
Рооја	67	147	12	43	52	0.31	144	256	25	94	81	0.36	0.78
Рооја	70	116	18	34	41	0.38	144	256	25	94	81	0.36	0.78
Gong	254	116	88	78	19	0.31	215	111	72	71	20	0.34	0.71
Gong	160	122	49	62	30	0.22	205	87	71	63	12	0.29	0.71
Gong	221	253	56	109	72	0.47	178	212	44	90	61	0.46	0.33
Gong	188	238	34	120	59	0.44	193	223	49	95	64	0.46	0.24
Zuo	556	504	160	236	134	0.48	628	628	142	344	142	0.5	0.02
rs1143634:													
Snoussi	428	182	157	114	34	0.29	306	94	120	66	14	0.24	0.24
Hefler	415	123	159	97	13	0.23	337	117	119	99	9	0.26	0.04
Balasubramanian	1062	320	410	242	39	0.23	629	211	231	167	22	0.25	0.24
Pooja	178	36	75	28	4	0.17	364	36	176	12	12	0.09	<0.01
Рооја	160	26	76	18	4	0.13	364	36	176	12	12	0.09	<0.01
Pooja	340	60	147	46	7	0.15	364	36	176	12	12	0.09	<0.01
rs1143627:													
Ito	219	235	58	103	66	0.48	155	215	28	99	58	0.42	0.18
Liu	379	351	102	175	88	0.48	579	683	133	313	185	0.46	0.98
Lee	344	362	96	152	105	0.49	285	295	70	145	75	0.49	0.99
Lee	209	203	51	107	48	0.58	211	219	43	125	47	0.49	0.02
Lee	553	565	147	259	153	0.49	496	514	113	270	122	0.49	0.12
Akisik	99	153	18	63	45	0.29	98	122	21	56	33	0.45	0.75
Gong	119	253	20	79	87	0.32	113	219	21	71	74	0.34	0.54
Gong	122	162	30	62	50	0.43	87	205	12	63	71	0.29	0.71
Gong	282	196	84	114	41	0.41	238	152	76	86	33	0.39	0.31
Gong	266	166	81	104	31	0.38	246	174	77	92	41	0.41	0.16
Zuo	510	550	136	238	156	0.48	630	626	144	342	142	0.49	0.03

contained 996 cases and 620 controls [13, 14]. Finaly, for rs1143627(C>T), just 2413 cases and 2438 controls were involved in five studies [10, 15–17, 19]. Genotype distributions in controls of all studies were in accordance with HWE.

Meta-analysis

The critical results of the current meta-analysis are described in Table III. Based on data, three common polymorphisms in IL-1 β (rs16944, rs1143634, rs1143627) were not significantly

Variable		Allele model			Nominant mode	-		ecessive mode		¥	pom snosvzom	-	Het		la
·	OR	(95% CI)	P-value ^a	OR	(95% CI)	<i>P</i> -value ^a	OR	(95% CI)	<i>P</i> -value ^a	OR	(95% CI)	P-value ^a	OR	(95% CI)	<i>P</i> -value ^a
rs16944:															
Overall	1.017	0.944-1.096	0.653	1.001	0.945-1.06	0.609	1.071	0.915-1.254	0.393	1.036	0.898-1.196	0.626	1.072	0.944-1.218	0.283
Ethnicity:															
Caucasion	0.986	0.837-1.16	0.862	0.97	0.863-1.09	0.813	1.029	0.703-1.507	0.082	0.988	0.657-1.487	0.956	1.039	0.847-1.273	0.714
Asian	1.071	0.98-1.172	0.13	1.01	0.927-1.101	0.436	1.229	1.063-1.422	0.005	1.083	0.884-1.327	0.443	1.217	1.06–1.396	0.005
African	1.017	0.927-1.104	0.791	1.043	0.938-1.161	0.979	0.937	0.763-1.15	0.533	1.041	0.889-1.219	0.615	0.901	0.743-1.092	0.289
rs1143634:															
Overall	1.072	0.783-1.466	0.665	1.036	0.777-1.382	0.807	1.27	0.863-1.87	0.226	1.277	0.758-2.151	0.359	1.243	0.86–1796	0.247
rs1143627:															
Overall	0.959	0.903-1.018	0.171	0.956	0.912-1.002	0.058	0.966	0.855-1.092	0.581	0.918	0.823-1.025	0.128	1.017	0.92-1.124	0.747
Ethnicity:															
Caucasion	0.918	0.726-1.159	0.472	0.94	0.792-1.115	0.475	0.879	0.611-1.265	0.488	0.879	0.617-1.252	0.473	0.943	0.779–1.143	0.553
Asian	0.965	0.899–1.034	0.312	0.944	0.897-0.994	0.027	1.023	0.881-1.187	0.768	0.929	0.812-1.064	0.289	1.084	0.963–1.22	0.181
African	0.988	0.873-1.118	0.849	1.024	0.924-1.135	0.651	0.866	0.632-1.186	0.37	0.932	0.689-1.259	0.645	0.845	0.639–1.118	0.239
Menopausal:															
Pre-menopausal	1.025	0.958-1.098	0.471	1.005	0.952-1.061	0.867	1.085	0.925-1.272	0.326	1.037	0.928-1.159	0.519	1.08	0.932-1.251	0.308
Post-menopausal	0.859	0.753-0.98	0.024	0.913	0.787-1.058	0.225	0.728	0.575-0.922	0.008	0.743	0.626-0.882	0.001	0.812	0.653-1.011	0.062

Table III. Summary ORs and 95% Cl of interleukin-1 β polymorphisms and breast cancer risk

 a Two-side χ^{2} test, P < 0.05 was considered statistically significant.

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Study	

ID	RR (95% CI)	% Weight
Caucasian		
Smith (2004)	0.66 (0.37, 1.18)	5.57
Hefler (2005) —	0.93 (0.58, 1.51)	7.21
Balasubramanian (2006)	0.95 (0.68, 1.34)	10.72
Gong (2013)	0.84 (0.46, 1.51)	5.35
Gong (2013)	→ 2.59 (1.38, 4.85)	4.88
Subtotal (<i>I</i> ² = 64.4%, <i>p</i> = 0.024)	1.03 (0.70, 1.51)	33.72
Asian		
Liu (2006)	1.35 (1.08, 1.70)	14.95
Pooja (2012)	1.20 (0.93, 1.55)	13.71
Pooja (2012)	1.09 (0.82, 1.45)	12.70
Subtotal (<i>I</i> ² = 0.0%, <i>p</i> = 0.483)	1.23 (1.06, 1.42)	41.36
African		
Gong (2013)	0.97 (0.73, 1.29)	12.71
Gong (2013)	0.90 (0.67, 1.21)	12.20
Subtotal (<i>I</i> ² = 0.0%, <i>p</i> = 0.718)	0.94 (0.76, 1.15)	24.92
Overall (<i>I</i> ² = 50.4%, <i>p</i> = 0.033)	1.07 (0.91, 1.25)	100.00
NOTE: Weights are from randorn effects analysis		
0.206	1 4.85	

Figure 1. OR of breast cancer in different ethnicities associated with rs16944 in IL-1 β gene for the TT genotype compared with the CC + CT genotype

Study ID	RR (95% CI)	% Weight
Caucasian		
Smith (2004)	0.77 (0.44, 1.34)	5.10
Hefler (2005)	1.06 (0.67, 1.67)	5.88
Balasubramanian (2006)	0.97 (0.70, 1.34)	12.30
Gong (2013)	0.89 (0.51, 1.56)	4.25
Gong (2013)	2.04 (1.12, 3.70)	2.72
Subtotal (/² = 38.6%, p = 0.164)	1.04 (0.85, 1.27)	30.24
Asian		
Liu (2006)	1.25 (1.01, 1.55)	20.40
Pooja (2012)	1.18 (0.93, 1.51)	11.73
Pooja (2012)	1.18 (0.91, 1.53)	10.00
Subtotal (<i>I</i> ² = 0.0%, <i>p</i> = 0.914)	1.22 (1.06, 1.40)	42.13
African		
Gong (2013)	0.98 (0.76, 1.28)	13.68
Gong (2013) –	0.82 (0.62, 1.09)	13.95
Subtotal (/² = 0.0%, p = 0.349)	0.90 (0.74, 1.09)	27.63
Overall (<i>l</i> ² = 34.7%, <i>p</i> = 0.130)	1.08 (0.97, 1.19)	100.00
0.27	1 3.7	

Figure 2. OR of breast cancer in different ethnicities associated with rs16944 in IL-1 β gene for the TT genotype compared with the CT genotype

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ID	RR (95% CI)	% Weight
Asian		
Ito (2002)	0.88 (0.80, 0.97)	11.17
Liu (2006)	0.91 (0.85, 0.98)	13.88
Lee (2006)	0.96 (0.88, 1.05)	11.92
Lee (2006)	0.95 (0.89, 1.02)	14.90
Akisik (2007)	1.06 (0.94, 1.19)	9.34
Subtotal (l ² = 42.1%, p = 0.141)	0.94 (0.90, 0.99)	61.20
Caucasian		
Gong (2013)	1.02 (0.95, 1.10)	13.77
Gong (2013)	0.86 (0.78, 0.95)	11.11
Subtotal (/2 = 86.8%, p = 0.006)	0.94 (0.79, 1.11)	24.88
African		
Gong (2013)	1.06 (0.92, 1.23)	6.95
Gong (2013)	0.99 (0.85, 1.14)	6.96
Subtotal (/² = 0.0%, p = 0.482)	1.02 (0.92, 1.14)	13.91
Overall (<i>I</i> ² = 53.2%, <i>p</i> = 0.029)	0.96 (0.91, 1.00)	100.00
NOTE: Weights are from randorn effects analysis		
0.779 1	1.28	

Figure 3. OR of breast cancer in different ethnicities associated with rs1143627 in IL-1 β gene for the CT + TT genotype compared with the CC genotype



Figure 4. OR of breast cancer in different menopausal state associated with rs1143627 in IL-1 β gene for the TT genotype compared with the CC + CT genotype



Figure 5. Begg's funnel plot analysis of publication bias. **A** – rs16944 in TT vs. CC + CT model. **B** – rs16944 in TT vs. CT model. **C** – rs1143627 in CT + TT vs. CC model. **D** – rs1143627 in TT vs. CC + CT model

associated with breast cancer risk in all genetic models. To go a step further, the data were stratified into different subgroups according to ethnicity, menopausal state, source of controls and genotyping methods. In terms of analysis by ethnic subgroup, TT genotype of rs16944 polymorphism represented a higher risk of breast cancer compared to CT genotype and CC + TT genotype and this was only significant in Asian people (TT vs. CC + CT: 1.229, 95% CI: 1.063–1.422, p = 0.005, Figure 1: TT vs. CT: 1.211. 95% CI: 1.057-1.388. p = 0.006, Figure 2), not in Caucasian and African populations. Similarly, it is only in Asian people that a significantly decreased breast cancer risk was found in the dominant model (CT + TT vs. CC: OR = 0.944, 95% CI: 0.897-0.994, p = 0.027, Figure 3) for the rs1143627 polymorphism. With patients being stratified according to the menopausal state, we noted that the rs1143627 polymorphism correlated with reduced breast cancer risk among post-menopausal women in three genotype models: the allele model (T vs. C: 0.859, 95% CI: 0.753–0.98, p = 0.024), the recessive model (TT vs. CC + CT: 0.727, 95% CI: 0.576-0.918. p = 0.007, Figure 4), and the homozygous model (TT vs. CC: 0.743, 95% CI: 0.626–0.882, *p* = 0.001). No significant associations between IL-1 β polymorphism and breast cancer risk were found in other stratified analyses based on source of controls and genotyping methods.

Publication bias

Egger's test and Begg's funnel plot were used so that we could evaluate the potential publication bias of the studied literature. No obvious evidence of publication bias was detected in IL-1 β (rs16944, rs1143634, rs1143627) (Figure 5). Also, good results were obtained in the sensitivity analysis (Figure 6).

Discussion

It is believed that cytokines are strongly connected to cancer pathogenesis accompanying increasing evidence, which suggests that they participate in tumor initiation, growth and metastasis [2]. Numerous studies about cytokine gene polymorphisms have been conducted to investigate their relations with many inflammatory and neoplastic diseases [14]. Emerging studies have reported the associations between IL-1 β polymorphisms and breast cancer risk because of IL-1 β 's

Meta-analysis estimates, given named study is omitted Α I Lower CI limit O Estimate | Upper CI limit Smith (2004) Hefler(2005) Liu (2006) Balasubramanian (2006) Pooja (2012) Pooja (2012) Gong (2013) Gong (2013) Gong (2013) Gong (2013) 0.91 0.98 1.09 1.21 1.25 Meta-analysis estimates, given named study is omitted B Lower CI limit OEstimate IUpper CI limit Smith (2004) 0 Hefler(2005) Liu (2006) Balasubramanian (2006) Pooja (2012) Pooja (2012) Gong (2013) Gong (2013) Gong (2013) Gong (2013) 0.92 0.97 1.19 1.08 1.24 Meta-analysis estimates, given named study is omitted С ILower CI limit OEstimate IUpper CI limit Ito (2002) Liu (2006) Lee (2006) Lee (2006) Akisik (2007) Gong (2013) Gong (2013) Gong (2013) Gong (2013) 0.91 0.92 0.95 0.98 1.00 Meta-analysis estimates, given named study is omitted D I Upper CI limit Llower CL limit OFstimate Ito (2002) Liu (2006) Lee (2006) Lee (2006) Akisik (2007) Gong (2013) Gong (2013) Gong (2013) Gong (2013) 0.84 0.89 1 1 2 0 97 1 0 6

> Figure 6. Sensitivity analysis. A – rs16944 in TT vs. CC + CT model. B – rs16944 in TT vs. CT model. C – rs1143627 in CT + TT vs. CC model. D – rs1143627 in TT vs. CC + CT model

crucial importance in breast cancer development. The size of samples in a single study was relatively small and the controls of some studies deviated from HWE, so the results were controversial. For example, Ito et al. first reported that rs1143627 was significantly associated with breast cancer risk [10] and another case-control study by Liu et al. verified this conclusion in the Chinese population [16]. However, Lee et al. and Akisik and Dalay did not find significant differences [15, 17]. Studies by Snoussi et al. and Pooja et al. revealed highly significant associations between rs1143634 and the aggressive phenotype of breast cancer [13, 18]. On the other hand, Hefler et al. and Balasubramanian et al. did not find significant differences [12, 14]. For rs16944, some studies indicated that rs16944 genotype reduced the risk of breast cancer, while others showed the opposite results.

Liu et al. conducted a meta-analysis to investigate the relations between three polymorphisms in IL-1 β gene and the risk of breast cancer [25]. The variant genotype of rs1143627 was found to be associated with a significantly increased breast cancer risk while the polymorphisms rs16944 and rs1143634 did not represent any associations with breast cancer risk, which was inconsistent with our results. As Wang et al. mentioned [26], the data reported by Liu et al. conflicted with the data from some previous publications, so the results provided by Liu et al. were untrustworthy. Given this situation, we conducted an updated meta-analysis to investigate the associations between three polymorphisms in the IL-1 β gene and breast cancer risk.

Some potential limitations in the present metaanalysis should be taken into consideration. First of all, compared to the one included in previous studies, this sample size was larger, but it was still relatively small for some SNPs and stratified analyses. Next, we could not conduct haplotype analysis and linkage disequilibrium, because more details about personal information regarding genotypes of the SNPs (rs16944, rs1143634 and rs1143627) in IL-1 β gene were unavailable. Furthermore, it is difficult for us to evaluate potential interactions between gene-environment, gene-gene and multiple polymorphic loci in the same gene. Regardless of these limitations, our present meta-analysis includes a much larger number of eligible studies and a stratified analysis.

In conclusion, our current meta-analysis suggests that the rs16944 polymorphism is significantly associated with increased risk of breast cancer among the Asian population, while the rs1143627 polymorphism reduces the breast cancer risk in post-menopausal women. In the future, we need further large-scale and rigorous studies to validate these findings.

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Conflict of interest

The authors declare no conflict of interest.

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